

1 **Title** Studying the spatial structuring of chemical elements through the prism of community and
2 landscape ecology

3
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43 The code required to replicate the analyses presented here is available on FigShare
44 (<https://figshare.com/s/64d154c3a8ec48784299>) while the data required to replicate the analysis is
45 available at Leroux, Shawn (2017): Stoichiometric distribution models: ecological stoichiometry
46 at the landscape extent. figshare. Dataset. <https://doi.org/10.6084/m9.figshare.4003482.v3>

47

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49 EH and SL designed the concept. All authors contributed to refining the framework. AM coded
50 the analysis to produce the results with contributions from EH, FM and NL. AM produced the
51 final figures. EH and AM wrote the first draft of the manuscript with contributions from SL. All
52 authors contributed to further editing of the manuscript. SL, EVW and YFW contributed the data
53 for the proof of concept.

54

55

56 **Abstract**

57 Approximatively 25 chemical elements are essential for the maintenance, growth and
58 reproduction of all living organisms. Hence, the movement, distribution, and relative proportions
59 of those elements on the landscape should influence the structure and functioning of biological
60 communities. Yet our basic understanding for the spatial distribution of elements across
61 landscapes is limited. Here, we propose to apply tools from community and landscape ecology to
62 study spatial patterns in elements. We illustrate this framework using tree leaves elemental
63 composition and demonstrate how spatial grain and spatial dissimilarity of elements interact
64 leading to predictable patterns in elemental distributions at various spatial scales. Meanwhile,
65 further analysis revealed that potassium and calcium are the most important elemental
66 contributors to spatial dissimilarity in leaf elements, raising new questions about their role in, or
67 response to, distributions of biodiversity and ecosystem functions. Our framework provides a
68 way to integrate abiotic and biotic processes, demonstrating how we can use community metrics
69 to investigate variability of individual elements across landscapes. We conclude by
70 hypothesizing that changes in the evenness or beta-diversity of elements should reflect the
71 structure of biotic communities, providing a long-sought mechanistic link between community
72 and ecosystem processes that can be measured directly in the field.

73

74 **Introduction**

75 Approximately 25 chemical elements are essential for the maintenance, growth and
76 reproduction of all living organisms (Elser *et al.* 1996; Kaspari & Powers 2016). The field of
77 ecological stoichiometry has made great progress in understanding how some elements (mainly
78 C, N and P) influence biotic community structure and stability through resource competition and
79 stoichiometric imbalance (Leroux & Schmitz 2015; Sterner & Elser 2017; Sentis *et al.* 2022).
80 Most of this research, however, has been conducted at small spatial extents (e.g., lab, mesocosm,
81 field sites; (Leroux *et al.* 2017). Consequently, our basic expectations for the spatial distribution
82 of elements and the drivers of those patterns at larger extents are limited to abiotic components
83 (e.g. bedrock; Kaspari & Powers 2016). To address this gap, researchers in marine (e.g.,
84 Galbraith & Martiny 2015), freshwater (e.g., Collins *et al.* 2017), and terrestrial (e.g., Leroux *et*
85 *al.* 2017) ecosystems are developing spatially explicit tools and methods for predicting spatial
86 patterns in elements at landscape/seascape extents.

87

88 Patterns of elemental abundance in geographic space are the result of the combined feedbacks of
89 passive abiotic flows and biotic ecosystem components (e.g., animal transport and deposition of
90 materials). Biotic influences on a landscape are expected to generate a patchy elemental
91 distribution characterized by areas with elemental hotspots and coldspots (Bernhardt *et al.* 2017;
92 McInturf *et al.* 2019). For example, McIntyre *et al.* (2008) demonstrated that the spatial
93 distribution and nutrient excretion of a diverse neotropical fish community can influence the
94 elemental distribution of N and P availability in streams and that these neotropical fish
95 communities can influence the distributions of N and P very differently. To scale up the study of
96 elemental distribution to regional extents, Leroux *et al.* (2017) proposed stoichiometric

97 distribution models (StDMs) – statistical models akin to species distribution models that make
98 spatially explicit predictions of elemental patterns. Stoichiometric distribution models generate
99 spatial maps of elements based on the measured relationship between elements and
100 environmental parameters (e.g., elevation, land cover), thus filling an important gap in measuring
101 elements across a whole landscape. Such elemental maps can then be used readily to study
102 spatial patterns of specific elements (e.g., percent C or N) or their ratios and can be linked to
103 patterns of resource quality (e.g., Heckford *et al.* 2022).

104

105 Building expectations on what spatial patterns in elements should look like, how they relate to
106 landscape features, and how they relate to the other biotic components of ecosystems is not
107 trivial. There is considerable variation in how different plant and animal species respond to and
108 consequently affect their environment – the scale of these effects is species-specific (see review
109 in Doherty & Driscoll 2018) and demonstrate hierarchical patterns within a species (e.g. Mayor
110 *et al.* 2009). For example, small mammals, such as snowshoe hares or meadow voles have a
111 relatively small home range and short daily movements (e.g., Rizzuto *et al.* 2021). Thus, these
112 smaller mammals interact with their environment at smaller scales than larger mammals (e.g.,
113 Leroux *et al.* 2017). The consequences of this smaller environmental footprint mean they may
114 contribute to more localized hotspots of nutrients. Larger mammals, such as moose, use an
115 environment several orders of magnitude larger than hares or voles in boreal forests (e.g.,
116 Balluffi-Fry *et al.* 2020). Through a larger home range and longer daily movements, these larger
117 animals can potentially create larger hotspots of nutrients or more dispersed nutrient hotspots
118 than animals with small home-ranges (e.g., around den sites; Gharajehdaghpour *et al.* 2016).
119 Movement is just one example of how a biotic component of an ecosystem may impact

120 elemental distributions across spatial scales, another example is demographic processes such as
121 the aggregation or over-dispersion of individuals. A tree (e.g., balsam fir) seedling may extract
122 water and limiting nutrients from a small area of soil surrounding its shallow roots whereas an
123 adult tree will extract these resources and affect a much larger area (Olesinski *et al.* 2011).
124 Alternatively, these plants and animals could be homogenizing resource distributions across their
125 home ranges by depleting hotspots that arise from abiotic influences.

126
127 The effects of these biotic ecosystem components on elemental distribution can have a temporal
128 component as well – foraging for resources whether by the root systems of plants or daily
129 movements of animals, are on a shorter temporal scale than dispersal and migration (e.g., Massol
130 *et al.* 2017). In this way a feedback develops both temporally and spatially, further complicating
131 spatial elemental patterns. For example, Pacific salmon spawning creates a pulse of nutrients
132 every fall (hotspot), which is subsequently used by plants and gradually depleted (cold spot) until
133 the next nutrient pulse (Helfield & Naiman 2001). Meanwhile, abiotic features of ecosystems can
134 also lead to patterns that depend on scale. At a broad spatial and temporal extent, weathering of
135 bedrock contributes to the elemental pool, which then get distributed via abiotic and biotic
136 mechanisms. For example, local hotspots of inorganic nitrogen in soils may develop beside
137 streams and rivers where ground and surface water mix, converging chemically distinct flow
138 paths (Edwards 1998). Human activities are significantly modifying both biotic (Díaz *et al.*
139 2019) and abiotic (Pörtner *et al.* 2022) components of ecosystems with potential knock-on
140 effects on the distribution and flux of elements (see Tucker *et al.* (2018) for examples of how
141 human activities have modified the movement of biotic ecosystem components). Consequently,
142 developing tools or methods to measure and quantify how spatial patterns in elements vary

143 across spatial grains is critical to diagnose how ecosystems are being shaped by the current era of
144 rapid change.

145

146 Here, we apply tools from community and landscape ecology to study spatial patterns in
147 elements (Fig. 1). We aim to explore how spatial grain and spatial dissimilarity of elements
148 interact, leading to predictable patterns in elemental distributions at various spatial scales. In
149 particular, because the distribution of elements across the landscape are the result of combined
150 *element-specific* feedbacks of passive abiotic flows and biotic ecosystem components, we
151 demonstrate how dissimilarity metrics offer a unique way to synthesize how multiple elements
152 vary independently across spatial gradients. We illustrate this framework with an empirical proof
153 of concept in the boreal forest. This proof of concept integrates models of the distribution of
154 elements in birch foliage across two landscapes (Leroux *et al.* 2017) with macro-ecological
155 smoothing techniques (Patrick & Yuan 2019) and dissimilarity analyses to investigate the
156 elemental diversity and variation across spatial grains. We first **a)** generate stoichiometric
157 distribution predictions for N, P, Ca, K, and Mg in birch foliage for our two landscapes. We then
158 **b)** use community ecology approaches to explore elemental dissimilarity in elemental content
159 across spatial scales, and **c)** identify potential elemental hotspots/coldspots including **d)** the
160 critical elements contributing to landscape uniqueness, i.e., hotspots or coldspots. Specifically,
161 we would expect dissimilarity to increase and asymptote at a landscape level dissimilarity value
162 as it is broadly driven by geological processes. We would expect hotspots/coldspots to become
163 more homogeneous across the landscape as spatial grain increases and geological processes
164 dominate biotic processes paralleled with a decrease in amount of core area for critical elements
165 contributing to landscape uniqueness.

166

167 **Using community and landscape ecology to study spatial patterns in elements: conceptual**
168 **underpinning**

169 We outline our empirical framework and proof of concept in three steps. We begin with a brief
170 description of our study system and stoichiometric distribution models (*sensu* Leroux *et al.* 2017)
171 as applied to this system. Next, we outline our application of tools from community ecology to
172 quantify different components of beta diversity applied to our proof of concept (Anderson *et al.*
173 2011; Legendre & De Cáceres 2013). Finally, we borrow metrics from landscape ecology (e.g.,
174 Wang *et al.* 2014) to measure how dissimilarity of elements varies across spatial grains in our
175 proof of concept.

176

177 *Proof of concept and stoichiometric distribution models*

178 For our proof of concept, we focused on two boreal forest landscapes (i.e., study areas), hereafter
179 called Plum Point and Old Man's Pond, on the island of Newfoundland, Canada (see
180 Supplementary Information 1, Fig A1 for map of study area locations). We clipped foliage
181 samples from the browse height available to boreal forest herbivores (i.e., leaves and small stems
182 from between 0.3 and 2 m) of 1 to 6 small trees (i.e., < 3 m) within 10 m radius plots across 106
183 plots at Plum Point between June 30 and July 7, 2015 (see Leroux *et al.* 2017; Balluffi-Fry *et al.*
184 2020 for more details on field sampling). We used birch trees (*Betula papyrifera*) < 3 m as part
185 of these data were originally designed to ask questions about moose foraging and moose only
186 forage on trees < 3 m. We expect these foliage samples to be indicative of bioavailable elements
187 for herbivores. In particular, white birch is the primary summer forage for moose and snowshoe
188 hares (Dodds 1960), some of the major herbivores in this system, and thus white birch foliage

189 serves as a proxy for bioavailable elements on the landscape. Focusing on this tree size, also
190 allows us to control for the possibility that trees of different sizes may differ in their elemental
191 composition. Foliage samples were pooled at the plot level and analysed for N, P, Ca, K, Mg (see
192 Leroux *et al.* 2017 for further details on sampling design and Supplementary Information 1 for
193 details on chemical analysis). We focused on these five elements because they are part of the
194 twenty-five elements essential for life and are among the most studied elements in ecological
195 stoichiometry (Jeyasingh *et al.* 2014), are critical for metabolic activity (Kaspari & Powers
196 2016), and frequently reported in studies of the ionome (or the mineral nutrient and trace element
197 composition of organisms; Parent *et al.* 2013). We omit carbon in our analyses because the
198 stoichiometric distribution models for Carbon had a poor fit (see Fig A2 and Table A1(a) for
199 model fit details). We focus only on N, P, Ca, K, and Mg, however, the approach presented here
200 could be applied to a broader suite of elements. Further, while we demonstrate the utility of this
201 approach in a focused proof of concept on the bioavailable elements in foliage samples, a similar
202 approach to the one presented here could be applied to elements in the soil, other important
203 forage, and other animals and would improve insights into the generality of the patterns observed
204 here. The emergence of synthetic databases of the elemental composition of soils and diverse
205 organisms (e.g., [https://stoichproject.wordpress.com/stoich/the-stoich-project/about-the-](https://stoichproject.wordpress.com/stoich/the-stoich-project/about-the-database/)
206 [database/](https://stoichproject.wordpress.com/stoich/the-stoich-project/about-the-database/)) may facilitate applications of our approach to different case studies.

207

208 Using this white birch foliage chemical composition and methods described in Leroux *et al.*
209 (2017), we fit StDMs to each element from the Plum Point study area. These StDMs are
210 statistical models that describe foliar elemental composition using environmental and biotic
211 covariates aimed at predicting correlations in organismal stoichiometry. We then extrapolate

212 elemental composition across Plum Point and Old Man's Pond landscapes from the correlations
213 in sampled sites of elemental composition with environmental covariates (Leroux *et al.* 2017).
214 Composition of white birch foliar proportions of mass of N, P, K, Ca, and Mg (i.e., elemental
215 composition; %) were the response variables. Predictor variables included three continuous
216 landscape covariates (i.e., normalized aspect or orientation of slope, slope and elevation), and
217 three categorical landscape covariates (i.e., landcover [coniferous, deciduous or mixedwood],
218 stand (canopy) height, and dominant tree species). For more details on StDM model fits, sample
219 design, and chemical analysis see Supplementary Information 1. We retained the full model for
220 each stoichiometric response to predict white birch across the landscape (R^2 for full models were
221 0.31, 0.42, 0.39, 0.13, and 0.27 for N, P, K, Mg, and Ca, respectively; see Table A1 and Fig. A3
222 – A7 in Supplementary Information 1 for specific details on StDM model fits). Using a raster for
223 each discrete covariate, we then developed stoichiometric composition predictions for foliage in
224 each cell. We then transformed the elemental compositions using a chord transformation
225 (Legendre & Gallagher 2001) based on the guidelines by Legendre & De Cáceres (2013) for
226 community composition data, of which our elemental composition data is akin (see Fig. A10 for
227 scaled elemental composition in the two study areas). In particular, this means that we can
228 calculate elemental dissimilarity using the techniques described in Legendre & De Cáceres
229 (2013) without violating any assumption of these approaches.

230 *Elemental dissimilarity across the landscape*

231 An important concept in community ecology is how species diversity is organized across space –
232 a concept for which many metrics have been developed including alpha and beta diversity. By
233 applying species diversity concepts to elemental composition across space, we can determine if
234 elemental composition varies across space in a predictable fashion. Despite many alpha-diversity

235 metrics incorporating both species richness and evenness, the constraints on both elemental
236 diversity and evenness mean that alpha-diversity is too coarse a metric to describe elemental
237 distributions appropriately across the landscape. As an explicit metric of variation, however,
238 Bray-Curtis dissimilarity (or variation in elemental composition between cells across the
239 landscape) is sensitive to deviations in evenness. Therefore, we applied dissimilarity metrics to
240 the transformed StDMs to examine stoichiometric variability across landscapes.

241
242 Before computing dissimilarities, we applied a macroecological spatial smoothing approach
243 (MESS), described in Patrick & Yuan (2019). This approach consists of applying a sliding
244 window across a landscape wherein each spatial window resampling and summarizing of local
245 observations occurs. To apply the MESS approach, a spatial grain (s) for the sampling regions is
246 selected, the number of random subsamples (n) of a given sample size (ss) of local cells within
247 that sampling region is specified, along with the minimum number (mn) of local cells in a given
248 sampling region (see Fig. 1). Using the inset in Fig. 1 as an example we would sample 6
249 subsamples (ss) of each window of size s and do this 3 times (n). By resampling the sub-samples
250 in this way we remove the impact of outliers (see Patrick & Yuan 2019). Similar to Patrick &
251 Yuan (2019), we employed a series of analyses to optimize the sample size (ss) and number of
252 random sub-samples (n) required to ensure sufficient sampling is done to determine the value of
253 the dissimilarity metric with appropriate accuracy (essentially where the changepoint in variation
254 around the Bray-Curtis dissimilarity metric occurred was deemed sufficient sampling; see Fig.
255 A8 and A9 in Supplementary Information 2 for results). Our optimal sample size was 11 ($mn =$
256 $1.5 \times ss$ as per Patrick & Yuan 2019), a n of 7, and a spatial grain increasing sequentially from a
257 grain size of 500 m to 5000 m in intervals of 200 m.

258 Dissimilarity metrics are then computed for each random subsample of the region, typically
259 averaged within each region. We calculated total community composition variance as proposed
260 by Legendre & De Cáceres (2013) as our dissimilarity metric which returns values between 0
261 (completely similar cells) and 1 (very dissimilar cells). In this way, we would expect Bray-Curtis
262 dissimilarity to be unaffected by grain size if there is no underlying pattern in elemental
263 distributions, increasing if there is an underlying gradient in elemental distributions, and
264 increasing much more rapidly if elements are aggregated in large patches (see Fig. 2a). Finally,
265 we compared the empirical patterns qualitatively to a null model by randomly shuffling
266 longitude and latitude couplets across elemental percentages and then re-running all analyses on
267 the random data set (100 iterations).

268

269 *Identifying hotspots and coldspots of elemental dissimilarity across the landscape*

270 We can extend our proof of concept to characterize cell uniqueness per study area in terms of
271 elemental contributions by applying the approach presented by Legendre & De Cáceres (2013)
272 and extended by Laliberté *et al.* (2020) (termed local contributions to beta diversity; LCBD).
273 This approach uses sum of squared deviations to characterize individual cells' contributions to
274 elemental beta diversity:

$$275 \quad LCBD = \frac{\sum_{j=1}^p s_{ij}}{SS_{Total}} \quad (1)$$

276 where

$$277 \quad s_{ij} = (y_{ij} - \bar{y}_j)^2 \quad (2)$$

278 and

$$279 \quad SS_{Total} = \sum_{i=1}^n \sum_{j=1}^p s_{ij} \quad (3)$$

280 Where i is cell and j is element, and y is elemental percentage. As such, this metric identifies the
281 cells that contribute more or less than the mean to beta diversity, acting as a comparative
282 indicator of cell uniqueness with larger LCBD values indicating cells that are more different in
283 terms of elemental composition. Importantly, this metric has been used to identify hotspots, or
284 cells which have unusually rich species compositions, and coldspots, or cells which are species
285 poor, but does not distinguish between the two (Legendre & De Cáceres 2013; Poisot *et al.* 2017;
286 Laliberté *et al.* 2020). We apply this approach to elements to identify unique cells on the
287 landscape – or cells that have unusually high elemental evenness and cells which have unusually
288 low elemental evenness.

289

290 *Identifying critical elements contributing to landscape uniqueness*

291 By applying the framework presented in Legendre & De Cáceres (2013) we can tease out site
292 uniqueness across the landscape, and can determine the individual contributions of each element
293 to each site across the landscape. In particular, we can identify patches where a specific element
294 is the most important contributor to elemental diversity (i.e., elemental contributions to beta
295 diversity; ECBD), and explore how the size of these patches change with grain size. For
296 example, if small foraging movements dominate the transport of elements around the landscape
297 at small grain sizes, we may expect a random configuration of very small patches for each
298 element, while if geochemical processes dominate the distribution of elements at larger grain
299 sizes then we may expect much larger patches for each element.

300

301 Similar to LCBD, our novel application we term, ECBD, uses a sum of squared deviations
302 approach:

303
$$ECBD = \frac{\sum_{i=1}^n s_{ij}}{SS_{Total}} \quad (4)$$

304 Where s_{ij} and SS_{total} are defined as above. To determine patchiness of elemental contributions, we
305 then applied landscape pattern metrics to determine how patches of elemental contributions vary
306 with grain size, and compared these values to the null model randomizations generated earlier in
307 the analysis.

308

309 We applied three spatial pattern metrics – mean core area, perimeter area fractal dimension, and
310 normalized landscape shape index – at the landscape level (see Fig. 2b for examples). These
311 metrics were identified by Wang *et al.* (2014) as suitable metrics to differentiate spatial
312 aggregations of landscapes.

313

314 Mean core area index is defined as the mean of the core areas of all patches belonging to each
315 landscape class (i.e. N, P, K, Ca, and Mg) where core area is defined as all cells that have no
316 neighbours (queen's case) with a landscape class different from their own. If geochemical
317 processes dominated the distribution patterns of elements at larger spatial grains, we would
318 expect the mean core area to increase with spatial grain whereas if mean core decreases with
319 spatial grain biotic ecosystem components (e.g., foraging) may be the drivers.

320

321 The perimeter area fractal dimension is a shape metric describing patch complexity, but is scale
322 independent, and has values between 1 and 2, with values closer to 1 denoting simple shapes and
323 values closer to 2 denoting irregular shapes. If biotic drivers (e.g., foraging movements) were
324 dominating the distribution pattern of elements at finer spatial grains, we may expect shapes with

325 higher overall irregularity (i.e., values closer to 2). However, abiotic drivers such as mixing of
326 surface and groundwater beside streams could also create higher overall irregularity.

327

328 The normalized landscape shape index is also an aggregation metric and describes the ratio of
329 the actual edge length of a landscape class to the hypothetical range of possible edge lengths for
330 this landscape class, although note that this landscape index was one of the metrics shown by
331 Frazier (2022) to be sensitive to variation in scope. This metric ranges from 0 if only one square
332 patch is present to 1 for a maximally disaggregated (i.e., checkerboard) landscape. If

333 geochemical processes were the only processes driving elemental distributions and they only
334 operated at larger spatial grains, then we would expect the normalized landscape shape index to
335 decrease from a value close to one to a value close to zero as spatial grain increases (Fig. 2b).

336 Meanwhile, if biotic ecosystem components such as demographic processes (e.g., aggregation
337 and over-dispersion) were the only processes driving elemental distributions we may expect the
338 normalized landscape shape index to increase from a value close to zero (clumping of species at
339 small grain sizes) to a value close to one at large grain sizes.

340

341 Finally, for these analyses, the pixel size is 6207.943 m² and the ratio of extent to grain (i.e.,
342 scope sensu Frazier 2022) is 187,154 and 205,842 for Plum Point and Old Man's Pond
343 respectively.

344

345

346 **A Community Ecology of Elements in Practice: Proof of concept results and discussion**

347 *Proof of concept and stoichiometric distribution models*

348 The white birch foliage StDMs explained between 13 and 42% of variation in the elemental
349 responses for N, P, K, Mg, and Ca (see Table A1 and Fig. A2 – A6 in Supplementary
350 Information 1 for specific details on StDM model fits). The estimated percent nitrogen ranged
351 from 1.90 to 3.67%, the percent phosphorus ranged from 0.16 to 0.42%, the percent potassium
352 ranged from 0.24 to 1.61%, the percent calcium ranged from 0 to 0.62%, and the percent
353 magnesium ranged from 0.03 to 0.24% for Plum Point. At Old Man’s Pond the estimated
354 percentages were a little bit higher for all elements except phosphorous. Nitrogen ranged from
355 1.78 to 3.72%, the percent phosphorus ranged from 0.13 to 0.43%, the percent potassium ranged
356 from 0.33 to 1.70%, the percent calcium ranged from 0.10 to 0.70%, and the percent magnesium
357 ranged from 0.11 to 0.24% (see Supplementary Information 3 Fig. A9 for scaled elemental
358 abundances in both study areas). These values are similar to patterns reported for other study
359 areas on the island of Newfoundland (Heckford *et al.* 2022).

360

361 *Elemental dissimilarity indices across the landscape*

362 In our proof of concept, we found that elemental dissimilarity increased with grain size,
363 eventually reaching an asymptotic “landscape level dissimilarity” when the grain size kept
364 increasing. The asymptote is the mean of the null model for our two study areas (Fig. 3). The
365 observed increase in dissimilarity with spatial grain mimics relationships already observed for
366 other components of biodiversity estimates. For example, species richness typically increases
367 rapidly at local scales due to a) stochastic variation in species occupancy patterns among cells
368 with more distant patches expected to demonstrate more differences in species composition (e.g.
369 gradient landscape of Fig. 2a); or b) increasing likelihood of sampling novel habitats as spatial
370 grain increases (e.g., Whittaker 1972; Harrison *et al.* 1992; Koleff & Gaston 2002). A variety of

371 different environmental and organismal factors are hypothesized to drive species dissimilarity at
372 various spatial scales ranging from habitat composition and structure at local scales to topology
373 and dispersal limitation at regional scales (see Table 1 in Barton *et al.* 2013 for more details).
374 Due to the way we carried out our analysis, it is unlikely that we are increasing the likelihood of
375 sampling novel elements as spatial grain increases since all elements were found at every site.
376 Rather, one explanation for what we are observing is a predictable gradient of variation in
377 elemental composition with more distant patches being more dissimilar: as the spatial grain of
378 the analysis increases, we observe an increase in dissimilarity as more dissimilar patches are
379 sampled (e.g., gradient in Fig. 2a vs. large checkerboard).

380

381 The trend we observe here between elemental dissimilarity and grain size suggests a similar
382 change in environmental and organismal drivers with scale. For example, organism movement
383 capacity affects species dissimilarity with species moving to track environmental gradients at
384 fine spatial scales (Kaspari *et al.* 2010). Thus, the rapid increase in elemental dissimilarity with
385 grain size could be related to consumer species distribution and activity including demographic
386 effects, abundance, and consumer movements as they distribute elements across the landscape.
387 Thus, the plateau at the asymptotic “landscape level dissimilarity” (Fig. 3) could be related to
388 differences between large patches driven by landscape features such as aquatic-terrestrial
389 boundaries, i.e., forest patch along the narrow lake on the southwest side of Old Man’s Pond (e.g.
390 McClain *et al.* 2003). Our example is only one testable hypothesis. Alternate hypotheses could
391 include variation in light availability, moisture, and plant competitors all of which could alter
392 local elemental composition potentially causing a similar increase in dissimilarity at finer grain
393 sizes.

394
395 Importantly, the dissimilarity in elements across the landscape can be a cause or a consequence
396 of biodiversity, but in either case, theory on nutrient co-limitations and stoichiometry would
397 suggest that greater elemental spatial turnover is indicative of greater biodiversity (Elser *et al.*
398 1996; Chase & Leibold 2009; Harpole *et al.* 2011). For example, if elemental dissimilarity is
399 very low with all elements equally abundant across the landscape, this should lead to biotic
400 homogenization and the dominance by a few species (ones with lower R^* , or equilibrium
401 resource level, for the specific homogenous condition). While our study presents some first steps
402 in tackling this question, further studies with species diversity data are needed to tease out these
403 directional relationships. However, recent work suggests that high imbalance in stoichiometric
404 ratios among connected patches (thus higher spatial turnover in elements) can lead to higher
405 productivity at the meta-ecosystem scale because of spatial complementarity in limiting nutrients
406 (Pichon *et al.* 2023).

407

408 *Identifying hotspots and coldspots of elemental dissimilarity across the landscape*

409 In our proof of concept, we demonstrate that at finer grains, the cells with higher local
410 contributions to landscape dissimilarity (i.e., those that more unique than the average site) cluster
411 together at both Plum Point and Old Man's Pond as evidenced by the yellow patches present at
412 grain size 500 m (Fig. 4). A grain size of 500 m corresponds to the scales at which organisms
413 feeding on birch leaves are most active in their foraging – for example snowshoe hares have a
414 home range of 0.027 - 0.042 km² (Rizzuto *et al.* 2021) while moose have been shown to have
415 daily movements between 0.5 km and 1.1 km depending on the season (Hundertmark 1998).
416 Whether these patches of high contributions of dissimilarity are a consequence of biotic

417 processes (e.g., foraging movements) or a cause of these biotic processes is unclear, and likely
418 are the result of a complex feedback between abiotic and biotic processes. Large-bodied
419 organisms can move an impressive amount of nutrients within and across ecosystems (Schmitz *et*
420 *al.* 2018), which can lead to the expected succession of hotspots and coldspots in the landscape
421 through source-sink dynamics (*sensu* Loreau *et al.* 2013 for the extension of the source-sink
422 concept to abiotic fluxes; McIntyre *et al.* 2008; McInturf *et al.* 2019). The clustering of cells with
423 high local contributions to landscape dissimilarity at finer grain sizes could be indicative of these
424 source-sink dynamics. Thus, the chain of cause and consequence generating the observed spatial
425 variation in resource elements is being reconsidered in an exciting junction between
426 biogeochemistry, ecosystem and community ecology.

427

428 As the grain size increases, however, the landscape becomes more homogenous with a greater
429 number of cells exhibiting an intermediate level of uniqueness, and those cells that are most
430 different are more diffuse across the landscape (Fig. 4). Only two larger patches of more unique
431 cells are observed at a grain size of 5000 m for Plum Point. Meanwhile the more unique cells in
432 both Plum Point and Old Man's Pond seem to be concentrated around bodies of water at a grain
433 size of 5000 m. One possible explanation for this is a shift in the driver of these
434 hotspots/coldspots from biotic to abiotic drivers, such as landscape features, for example aquatic-
435 terrestrial boundaries. Likely, it represents a complex feedback between abiotic drivers, such as
436 the mixing of surface and ground waters at the aquatic-terrestrial boundary converging
437 chemically distinct flow paths (Edwards 1998), and biotic drivers, such as terrestrial ungulates
438 foraging on aquatic plants (e.g. moose-aquatic plants relationship; Bump *et al.* 2017). Future
439 work can apply our method to more completely account for the stocks and flows of elements

440 across scales including measuring the elemental footprints (or combination of nutrient stocks and
441 effects of organisms on stocks) of soils, plants and small and large herbivores.

442

443 *Identifying critical elements contributing to landscape uniqueness*

444 In our proof of concept, we observed that in both Plum Point and Old Man's Pond, potassium (K)
445 and calcium (Ca) are the most important contributors to dissimilarity. This is intriguing as many
446 studies to date have been focused on the contributions of the three more abundant elements –
447 Carbon, Nitrogen, and Phosphorus to biotic community structure and stability (e.g., Leroux &
448 Schmitz 2015; Sterner & Elser 2017; Sentis *et al.* 2022). Rather what this result suggests is that
449 the distribution of these less abundant, but essential, elements may be just as important for
450 community structure. Specifically, in both study areas K is more often the most important
451 contributor (see Supplementary Information 3 Fig. A11; Fig. 5). In Plum Point, however, Ca
452 becomes increasingly important as grain size increases, while K becomes less important. The
453 converse occurs in Old Man's Pond where K becomes increasingly more important, but appears
454 to level off at 4000 m. Interestingly, Ca and K are inherently connected in plant nutrition
455 whereby as Ca concentrations decrease, K concentrations increases due to luxury consumption of
456 K and K antagonism whereby K occupies the majority of exchangeable ionic sites (Wilkinson *et*
457 *al.* 2000). Moreover, it has been shown that Ca is generally taken up in amounts corresponding
458 to availability rather than plant requirements (Knecht & Göransson 2004), perhaps explaining the
459 different patterns in the two study areas. Our results highlight the importance of rarely measured
460 elements such as K and Ca for structuring spatial dissimilarity in elements across spatial grains, a
461 result that underscores the difficulties in multi-dimensional ratios given the controversy in ratio
462 approaches for more elements than just N, P, and C (Parent *et al.* 2013).

463
464 There is a decline in mean core area with grain size for both elements in both study areas. The
465 decline is especially pronounced for K in Plum Point (Fig. 5). The decline demonstrates that the
466 landscape gets more homogeneous as grain size increases and occurs most precipitously for Ca
467 in both Plum Point and Old Man's Pond and then levels out at a grain size of 1500 m for both
468 study areas. These declining trends are reflected in the normalized shape index plots where both
469 elements become less aggregated as grain size increases. However, in the Plum Point study area
470 there is a peak for both K and Ca at a grain size of 2000 m, after which the normalized landscape
471 shape index declines. Elements at Plum Point are therefore least aggregated at grain size of 2000
472 m, and the landscape becomes more like a single square patch as grain size increases. At Old
473 Man's Pond the normalized shape index reaches an asymptote for both K and Ca at a grain size
474 of ~3000 m. Examining the perimeter area fractal dimension illustrates that, the shapes of
475 elemental patches across the study area become more irregular as grain size increases, contrary
476 to our earlier speculations. One explanation for this could be the increasing importance of abiotic
477 drivers such as confluence of ground and surface waters at aquatic boundaries at larger grain
478 sizes. Grain size contributes to increasing perimeter area fractal dimension irregularity in both
479 study areas, but does appear to asymptote at a grain size of approximately 2000 m for Ca in both
480 study areas. However, it should be noted that the two study areas have a different scope (ratio of
481 range to extent) and Frazier (2022) showed that some metrics, including the landscape shape
482 index, which we used here, is highly sensitive to variation in scope. Thus, comparisons between
483 the two study areas may be artifacts of differences in scope.
484

485 Much research has determined that elements cluster in two distinct groups based on the
486 competing needs to grow and maintain existing structure (Ågren & Weih 2012; Zhang *et al.*
487 2012). The first reflects plant growth rate and correlates with nitrogen and phosphorus
488 concentrations – elements critical for the metabolism of nucleic acids and proteins (Ågren &
489 Weih 2012). The second reflects maintaining plant structural components requires K, Ca, and
490 Mg, that are critical for cell walls and might be more important for long-lived slow-growing
491 plants (Nakajima *et al.* 1981; Maathuis 2009). The structural elements appeared to be the most
492 important drivers of elemental dissimilarity in white birch tissue across the landscape in our
493 proof of concept. The ECBD metric does not tell us whether elemental importance is caused by
494 low amounts or high amounts of these structural elements, however evidence suggests that K is
495 rarely limiting and excess K availability can contribute to higher growth rates in plants (Maathuis
496 2009). Similarly, Ca is usually abundant in the lithosphere, however, Ca concentrations in plants
497 can fall below a critical threshold in fast-growing tissues which can lead to diseases such as
498 blossom end rot in tomatoes (Maathuis 2009).

499

500 Notwithstanding current challenges in identifying the processes behind elemental importance,
501 the fact remains that we can observe patterns in elemental importance. In meadow systems,
502 studies have shown that in the absence of grazing by sheep, plants become more nutritious with
503 greater concentrations of micronutrients such as Ca (Marrs *et al.* 2020). In the boreal forest,
504 moose are selective herbivores and could be selectively removing nutritious plants that relatively
505 high in N and P in a similar fashion creating elemental coldspots (Pastor *et al.* 1998). Indeed, the
506 observed patterns in elemental importance should have a predictive influence over biodiversity
507 simply because these patterns control the amount of co-limiting nutrients and their spatial

508 turnover (Marleau *et al.* 2015). More specifically, the spatial heterogeneity in resource elements
509 should relate to coexistence dynamics and community structure directly influencing the number
510 of available niches. For example, studies have shown a correlation between plots containing tree
511 species with calcium-rich detritus and a greater diversity and biomass of earthworms (Reich *et*
512 *al.* 2005).

513

514 **Perspectives**

515 Our understanding of how and why chemical elemental concentrations change in space or across
516 scales is currently very poor (Kaspari & Powers 2016), contributing to poor integration of
517 empirical and theoretical metaecosystem research. The spatial distribution of resource elements,
518 however, reflects element-specific feedbacks of passive abiotic and biotic processes, thus
519 dissimilarity in resource elements may contain much of the information necessary to infer key
520 characteristics of biotic communities, and the expected relationship between those characteristics
521 and ecosystem function (Wu *et al.* 2015). Therefore, identifying spatial patterns in resource
522 elements in the landscape may allow regional scale prediction of community structure,
523 biodiversity, and ecosystem function. Moreover, the analysis of spatial patterns in the
524 fundamental building blocks of all life – elements – can pave the way for greater integration of
525 empirical and theoretical metaecosystem research. Such an integration would open significant
526 new opportunities to finally develop a unified field of spatial ecology, from community to
527 ecosystem and landscape level processes. Disentangling the drivers of these patterns is
528 challenging, especially as the direction of causality is not always intuitive and often involves
529 feedbacks between abiotic and biotic ecosystem components (see Ellis-Soto *et al.* 2021 for a
530 guide to measuring the biotic component of spatial ecosystem subsidies). Here, we have

531 presented a framework to explore spatial patterns in elemental distributions from foliage
532 samples; however, the approach we present is timely as it corresponds with a time of recent
533 growth in data on elemental composition of soil and diverse organisms across many different
534 biomes. In this way, we provide ecologists with an additional and critical tool for describing
535 ecosystems (see Table 1 for explicit predictions).

536

537 Dissimilarity of resource elements may be the result of or may affect food web structure. For
538 instance, we know that elemental hotspots tend to appear at locations in the landscape rich in
539 energy and material exchange (e.g., deep-sea hydrothermal vents), or where different chemical
540 reactants are expected to meet (e.g., ecotone between terrestrial and aquatic ecosystems -
541 McClain *et al.* 2003; Bernhardt *et al.* 2017). Different lines of evidence also suggest that
542 organism movement can track those hotspots in the landscape (McNaughton *et al.* 1988;
543 Leimgruber *et al.* 2001; Leroux *et al.* 2017; Balluffi-Fry *et al.* 2020; Kaspari 2020). Recent
544 conceptual and theoretical developments related to metaecosystem theory now challenge our
545 current understanding by suggesting that organisms themselves, through their movement in the
546 landscape and feeding interactions, could generate the emergence of those observed elemental
547 hotspots in the landscape (i.e., zoogeochemistry; Leroux *et al.* 2017; Gounand *et al.* 2018;
548 Schmitz *et al.* 2018; McInturf *et al.* 2019) and potentially diffuse the elemental hotspots
549 generated by abiotic processes. The elemental dissimilarity approach presented here provides a
550 way of measuring some of these elemental patterns across the landscape with the ECBD metric
551 being particularly useful in teasing out the scale at which each element is most important for
552 dissimilarity (see Table 1 for testable hypotheses).

553

554 More than that however, previous research has demonstrated how energy and nutrients in food
555 webs generally come from different trophic pathways that vary in their resource quality and the
556 speed at which energy and materials flow (Rooney *et al.* 2006). Traditionally, stoichiometric
557 analyses focus on ratios, which have proven useful to understand trophic interactions (e.g., Hall
558 2009; Leroux & Schmitz 2015) and co-limitation in competitive communities (e.g., Harpole *et*
559 *al.* 2011; Marleau *et al.* 2015). In particular, these stoichiometric analyses that have focused on
560 the C:N and C:P ratios of resources are a useful way to characterize these different trophic
561 pathways as ecosystems with resources with high (low) C:N ratios tend to be less (more)
562 palatable for herbivores and this decreases (increases) the rate of materials flowing in this
563 pathway (Hall *et al.* 2004; Shurin *et al.* 2006). Our work, however, has demonstrated the role
564 that some of the other less abundant essential elements (e.g. K and Ca) play in landscape
565 elemental dissimilarity – potentially offering additional currencies to tease out some of these
566 relationships. For example, whereas some top consumers focus on one energetic or material
567 pathway, many consumers couple multiple pathways (Rooney *et al.* 2008; Ward *et al.* 2015).
568 Indeed, this coupling of different energetic pathways is an important contributor to ecosystem
569 stability (McCann *et al.* 1998; Rooney *et al.* 2006). We hypothesize that, irrespective of the
570 direction of causality, changes in the evenness or dissimilarity of element ratios should be
571 reflected directly into the structure of biotic communities, providing a long-sought mechanistic
572 link between community and ecosystem processes that can be measured directly in the field.
573 Thus, a community perspective on elemental resources holds promise to synthesize our
574 understanding of biodiversity and ecosystem function across level of organization.

575

576 From an applied perspective, the majority of conservation planning focuses on the protection and
577 restoration of species, populations or communities, despite a call towards “ecosystem-based
578 management” in terrestrial, freshwater, and marine realms. We surmise that, in some cases,
579 elemental hotspots may be important features to protect or restore. For example, natural sodium
580 licks are critical sources of this limiting nutrient for many large ungulate communities in
581 temperate and boreal forest ecosystems (Kaspari 2020). In many ways, a pivot towards
582 considering elemental hotspots for conservation is akin to calls for considering species
583 interactions (Tylianakis *et al.* 2010; Harvey *et al.* 2017) – i.e., levels of biodiversity beyond the
584 species that may be critical to maintain ecosystem functioning. We provide an analytical pipeline
585 to begin identifying elemental hotspots across landscapes. Equally, however, we provide an
586 analytical pipeline with the potential to identify areas to protect. For example, our analyses
587 suggest several areas that stand out as hot/coldspots of elemental dissimilarity (e.g., the yellow
588 patches present at grain size 500 m in Fig. 4). If future research demonstrates that these patterns
589 in elemental dissimilarity are reflected in species diversity as we have hypothesized, then our
590 analytical approach provides a tool to identify the specific areas (including area boundaries) to
591 protect.

592

593 Geological processes and abiotic factors have historically been the point of focus for predicting
594 elemental concentrations at local study sites, but recent works have shed light on the significance
595 of biotic ecosystem components on elemental distribution (Gounand *et al.* 2018; Schmitz *et al.*
596 2018; McInturf *et al.* 2019; Schmitz & Leroux 2020; Malhi *et al.* 2022). Biotic and abiotic
597 processes affecting the distribution of elements are often interdependent. By integrating both
598 perspectives, meta-ecosystem theory illustrates the intimate feedback between the biotic and

599 abiotic components of ecosystems, and how patch-specific biotic (e.g., organismal dispersal) and
600 abiotic (e.g., inorganic nutrient runoff) processes can lead to the emergence of regional scale
601 phenomena (Loreau *et al.* 2003; Massol *et al.* 2011, 2017). However, beyond theoretical and
602 conceptual advancements, we are still missing a coherent empirical framework to i) study and
603 analyse empirical patterns of elemental distribution in space and across spatial grains, and ii) to
604 link those abiotic observations (including information on longer timescale abiotic drivers such as
605 weathering of parent geological material) to the biotic component of ecosystems. We are
606 optimistic that our community perspective on elemental resources helps develop an empirical
607 framework initially to enhance our descriptions of ecosystems and understanding of how
608 elemental distributions scale with spatial gradients across different ecosystem types (see Table
609 1). We feel that our work provides the foundation upon which future work can attempt to address
610 (ii) - linking the abiotic observations to the biotic components, thus truly bridging these distinct
611 ecosystem compartments.

612

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617

618

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807 **Tables**

808 **Table 1.** Testable hypotheses generated from our proof of concept.

Metric	Hypothesis	Example of data
Elemental	Spatial grain at which the change in	Elemental data of forage
Dissimilarity	elemental dissimilarity is the greatest is also the grain size at which organisms are most active in their foraging	species Telemetry data for consumer species
	Dissimilarity in resource elements follow general patterns with spatial scale	Elemental data of soil Elemental data of individual forage species
	Changes in dissimilarity of elements with spatial scale should be reflected in the structure of biotic communities. Specifically:	Elemental data of soil Species diversity data
	1. Low spatial turnover in resource elements should lead to the assembly of functionally more similar species, a lower alpha- diversity, and a generally lower functioning level (i.e. saturation is	

reached faster because of functional redundancy amongst species).

2. High spatial turnover in elements, on the other hand, should lead to more functional complementarity and generally higher functioning level. In particular, lower local evenness of elements should lead to dominance by a few species that are able to maximize their use of most elements being at very low level (R^*).

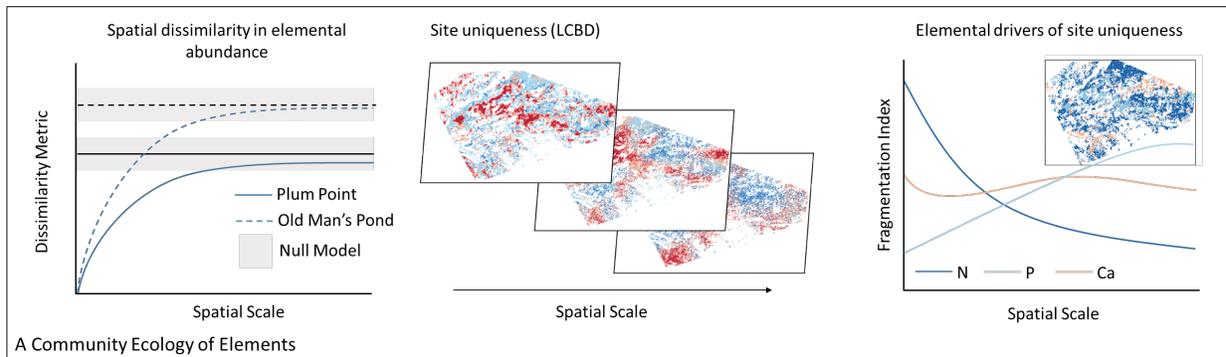
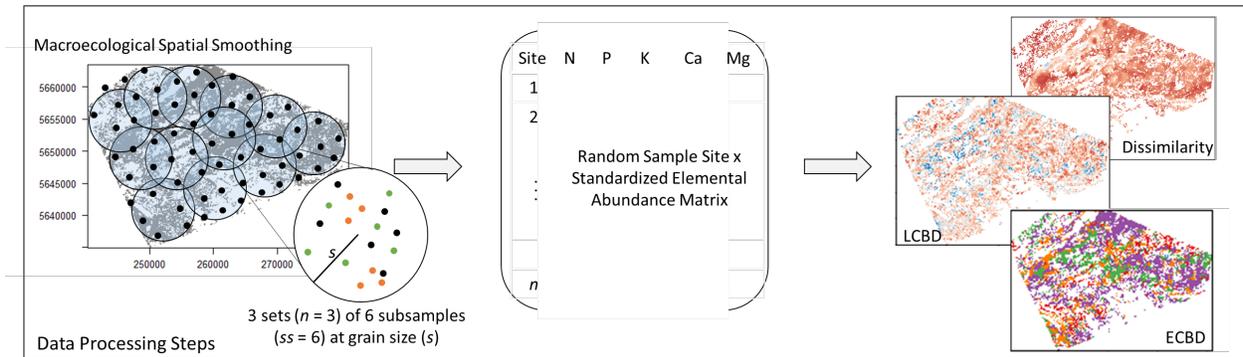
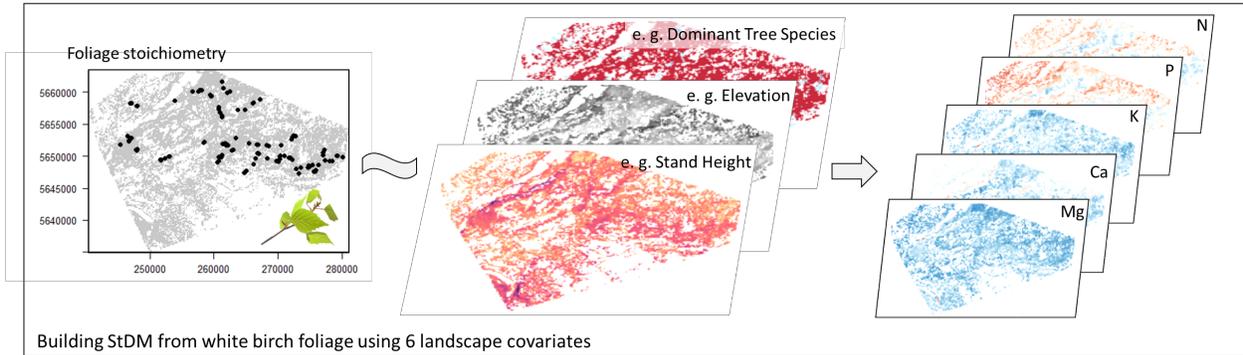
Local contribution to elemental dissimilarity	Localized hot and cold spots of elemental dissimilarity should reflect localized hot and cold spots of species diversity at finer grain sizes and hot and cold spots of landscape diversity (i.e. aquatic-terrestrial boundaries) at large grain sizes	Diversity surveys Satellite imagery
	Alternatively, local hot and cold spots of elemental dissimilarity could emerge through the movement of organisms across the landscape	Telemetry Theoretical models

Elemental contribution to dissimilarity	Elements may differentially indicate areas of high biodiversity - for example tree species with calcium-rich detritus support greater diversity and biomass of earthworms (Reich et al. 2005).	Biodiversity surveys in areas where calcium is the most important element contributing to elemental diversity
	Alternatively, selective herbivory for limiting nutrients could lead to elemental cold spots	Herbivory surveys in areas where calcium is depleted

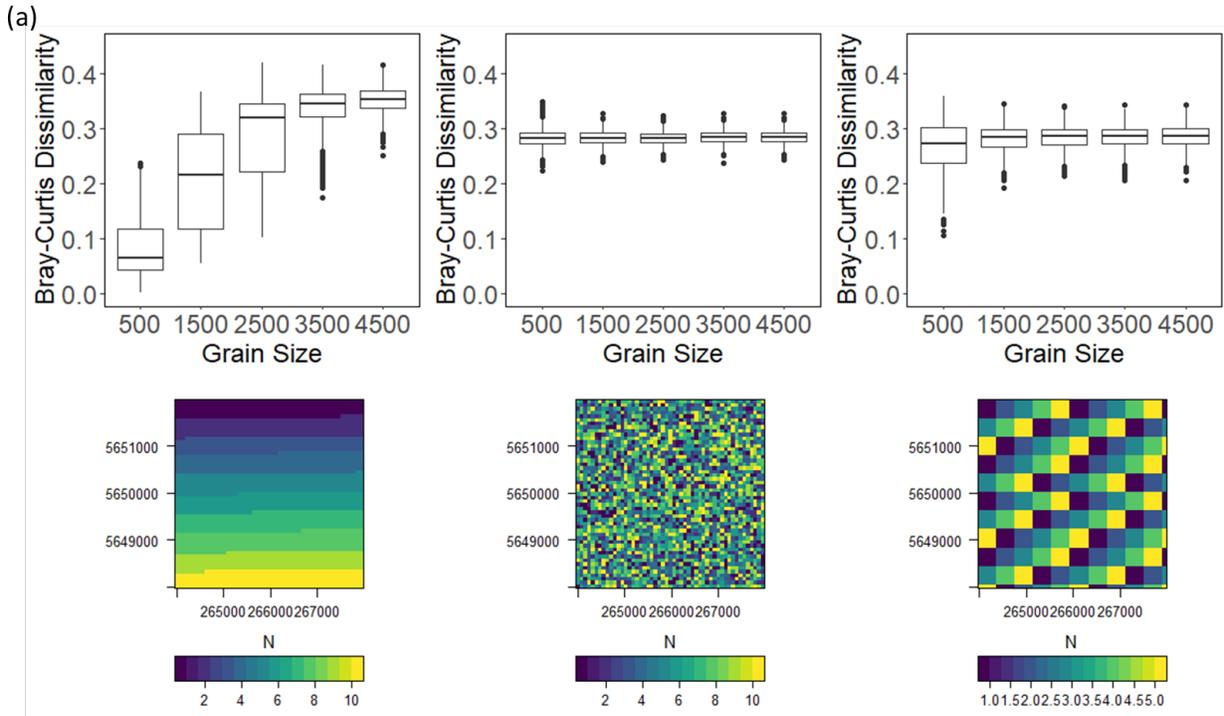
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811 **Figures**

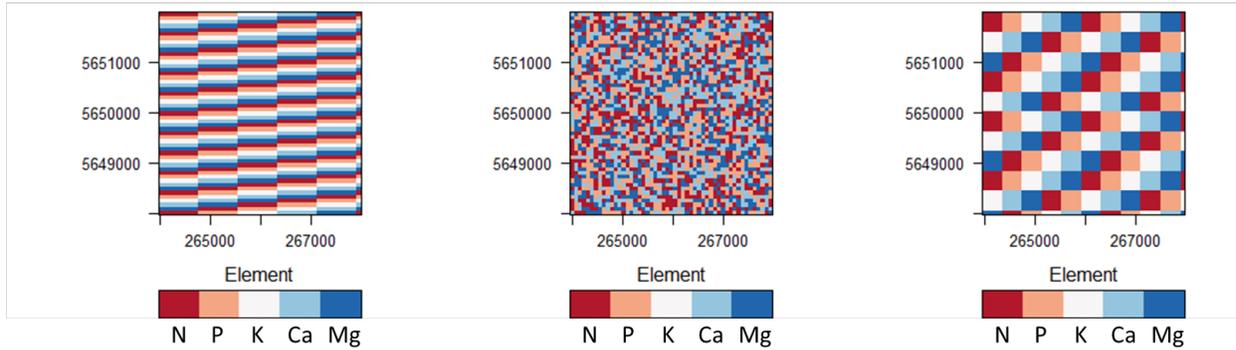


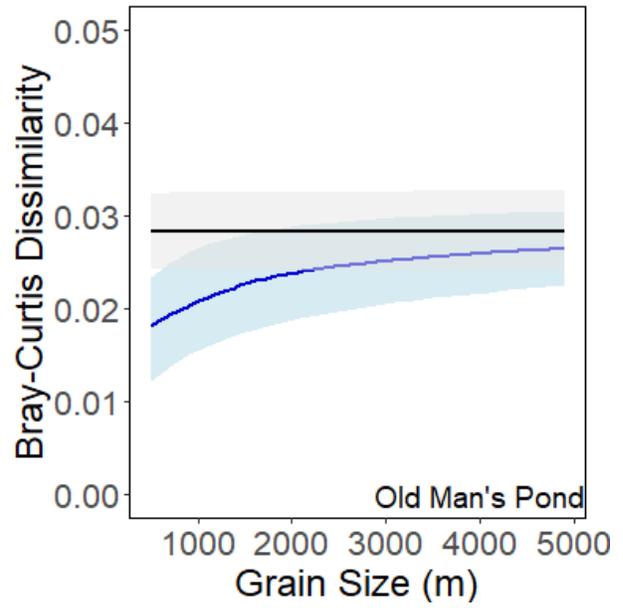
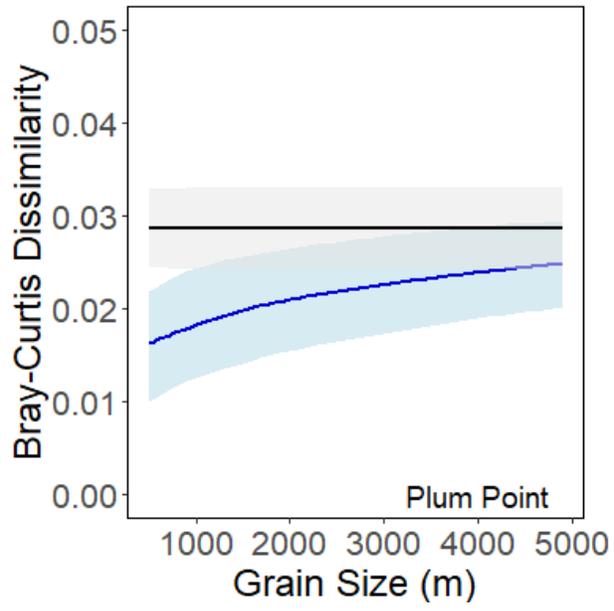
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(b)

	Patchy Landscape		Random Landscape		Checkerboard Landscape	
	N	Mg	N	Mg	N	Mg
Mean Core Area	0	0	0.01	0.02	5.08	5.32
Perimeter Area Fractal Dimension	1.75	1.93	1.70	1.73	1.03	1.14
Normalized Landscape Shape Index	0.54	0.54	0.74	0.75	0.19	0.18

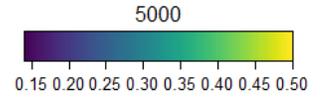
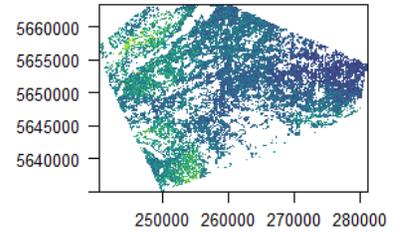
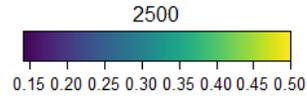
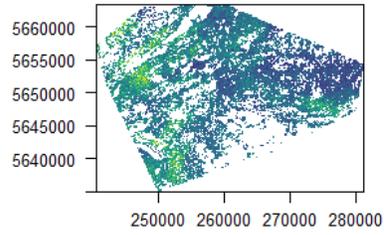
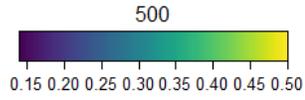
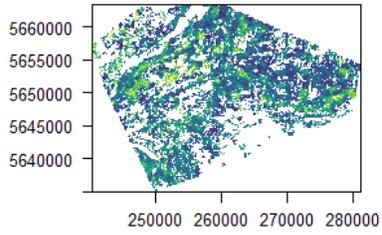




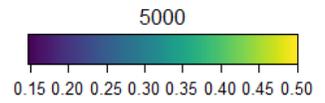
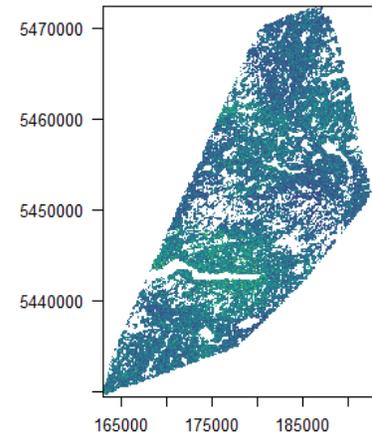
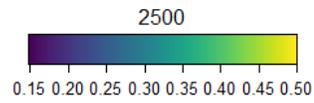
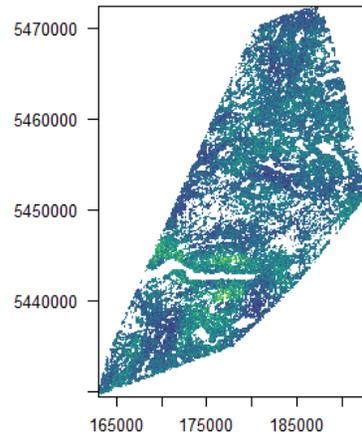
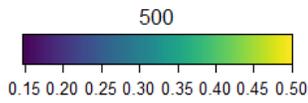
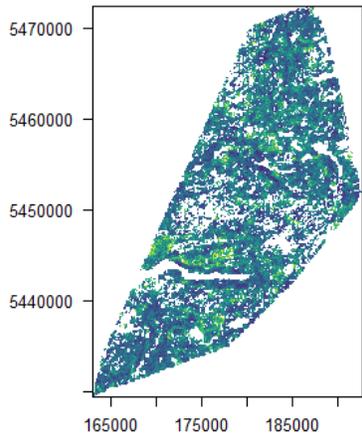
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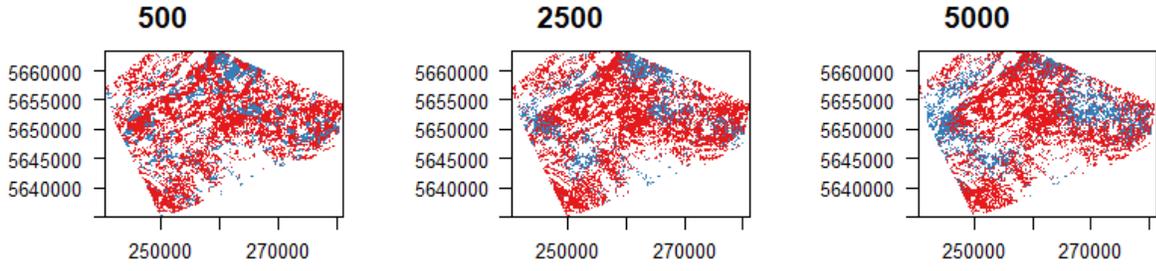
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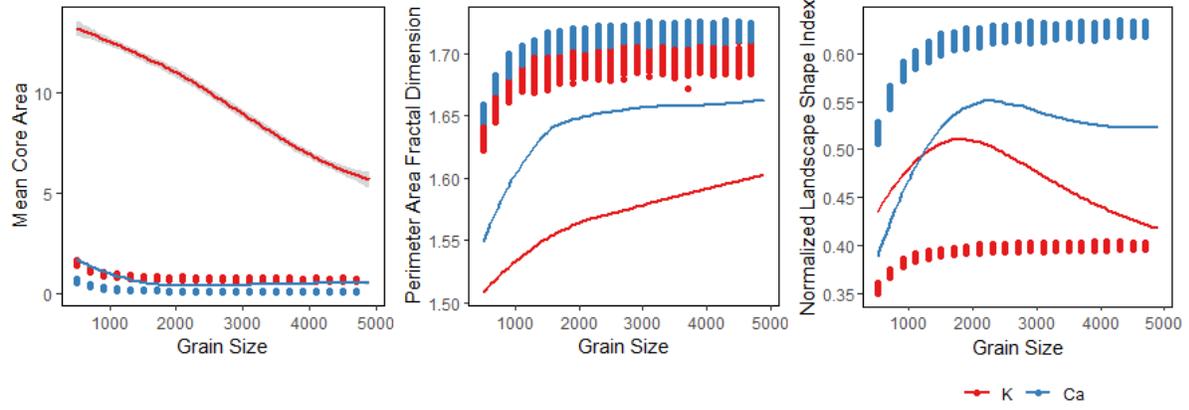
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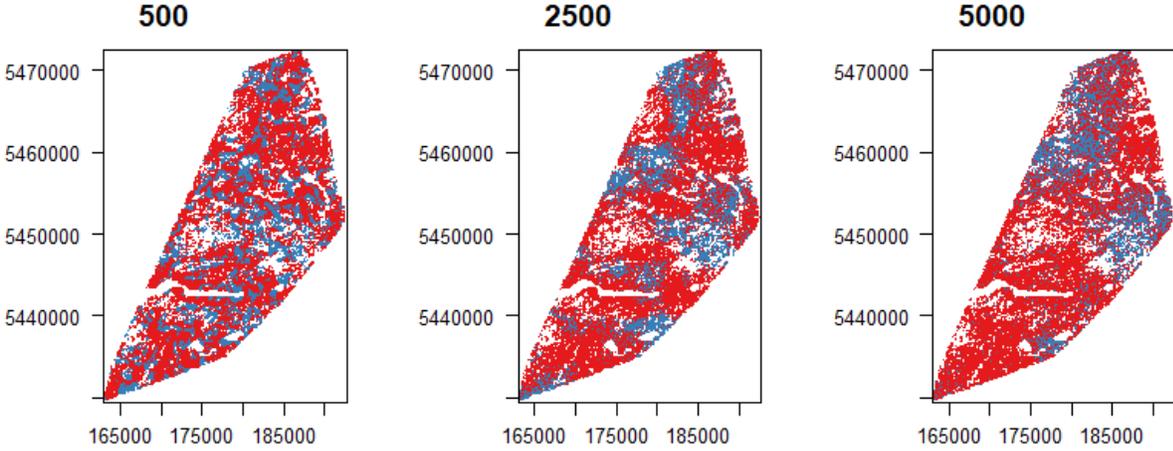


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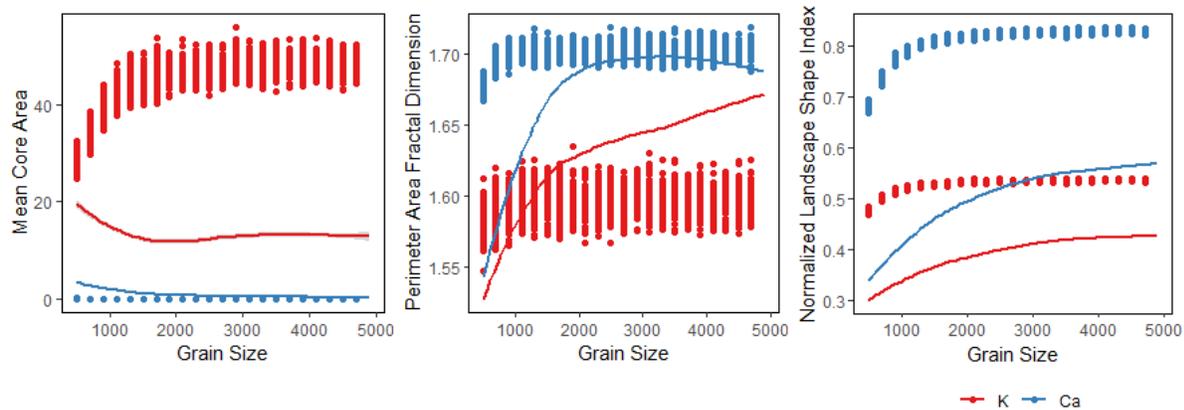


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823 (a)



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826 (b)

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828 **Figure Captions**

829 **Figure 1.** The empirical framework can be outlined in three steps. First, we build a
 830 stoichiometric distribution model (sensu Leroux et al. 2017) for each of our five key elements
 831 from white birch foliage using three landscape covariates (normalized aspect, slope, and
 832 elevation) and three categorical landscape covariates (landcover, stand height, and dominant tree
 833 species) to predict stoichiometric composition for foliage in each cell. Then, we apply the macro-
 834 ecological spatial smoothing approach proposed by Patrick and Yuan (2019) to resample and
 835 summarize local observations at increasingly large grain sizes to determine how community
 836 metrics (dissimilarity, local contribution to dissimilarity, and elemental contribution to
 837 dissimilarity) applied to elemental composition varied with spatial scale. We apply this approach
 838 to both the landscapes and a randomization, null model, of each landscape. Finally, we report
 839 how our three community metrics - dissimilarity in elemental composition, local contribution to
 840 dissimilarity, and elemental contribution to dissimilarity – vary with spatial scale.

841 **Figure 2.** Demonstration of how (a) dissimilarity changes with grain sizes for different
 842 hypothetical landscapes – one in which elemental composition exists in a predictable gradient,

843 another randomized landscape, and a third where elemental composition is a checkerboard of
844 high composition and low composition. Demonstration of how (b) patch metrics change with
845 different hypothetical patch configurations – one where patches are organized rectilinearly on the
846 landscape, one where patches are randomly scattered across the landscape, and a third where the
847 patches are an even checkerboard. Recall patch class is based on which of the five elements has
848 been ranked most important contributor to pixel dissimilarity. In both (a) and (b) the numbers on
849 the x- and y- axes are in UTM coordinates.

850 **Figure 3.** Relationship between elemental dissimilarity indices and grain size for Plum Point and
851 Old Man’s Pond. Here, the observed data is presented in blue with the mean value in dark blue
852 and the light blue indicating the standard deviation, while the black line indicates the mean value
853 of null model simulations (where the location of each site is randomly moved in the landscape)
854 with the light grey indicating the standard deviation.

855 **Figure 4.** LCBD across the two landscapes (Plum Point top, Old Man’s Pond bottom) at three
856 different grain sizes where the x- and y-axes are in UTM coordinates. Here, blue indicate sites
857 that contribute more than average to beta-diversity, while red indicates sites that contribute less
858 than average to beta-diversity. Note that a site can contribute more than average to beta-diversity
859 by being either more nutrient-rich or nutrient-poor than the average site. See Fig. A1 in
860 Supplementary Information 1 for map of study area locations.

861 **Figure 5.** Landscape plotted by which element contributes the most to its beta diversity for three
862 representative grain sizes (500, 2500, and 5000 m) for (a) Plum Point and (b) Old Man’s Pond
863 where red is K and blue is Ca and white cells are areas without forests (e.g., lakes, rivers, roads).
864 Bottom row in both a) and b) illustrates of how landscape metrics measuring how patches with
865 the same top elemental driver of beta diversity (ECBD) change with grain size for each element.

866 Here the lines indicate the observed data while the points are the results from the null model. See
867 Fig. A1 in Supplementary Information 1 for map of study area locations.

868